



# Invariant Disns for Markov Chains contd

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March 29, 2021

## Irreducible periodic case

This argument used explicitly aperiodicity. Indeed its conclusion is not true for periodic chains. Nonetheless it has consequences for periodic irreducible chain. We know that an irreducible chain has a single period  $d$ . If the chain is periodic, then  $d > 1$ . We suppose that to be the case. In that case for  $j \in I$  we can define  $D_0$  to be the set of sites so that for some  $n$ ,  $p_{jk}^{nd} > 0$ . So equivalently,  $k \in D_0$  if and only if  $p_{jk}^{nd} > 0$  for all  $n$  large. We can then define for  $0 \leq r < d$ ,  $D_r = \{k : p_{jk}^{nd+r} > 0\}$  for some  $n = \{k : p_{jk}^{nd+r} > 0\}$  for all  $n$  large.

## Irreducible periodic case

It is easily seen that if  $i \in D_s$ ,  $k \in D_t$ , then  $p_{ik}^n > 0$  only if  $s + n$  is congruent to  $t \bmod(d)$ .

Thus starting at, say,  $j$  the chain moves from  $D_0$  to  $D_1$  to  $D_2$  and so on  $\bmod(d)$ . The transition matrix  $P^d$  is irreducible on  $D_0$ , on  $D_1$  . . but not on  $I$ .

### Theorem

*If  $\pi$  is invariant distribution for periodic irreducible Markov chain then for  $D_r$  as above, for fixed  $r$*

$$\lim p_{jk}^{nd+r} = d\pi(k)$$

*for every  $k \in D_r$*

Remark: Obviously if  $k$  is not in  $D_r$ , then for every  $n$ ,  $p_{jk}^{nd+r} = 0$ .

# Reversible Processes

An important class of Markov chains are so-called *Reversible* processes: A Markov chain is reversible with respect to a measure  $\pi$  if

$\forall x, y \in I, \pi(x)p(x, y) = \pi(y)p(y, x)$ . Most chains are not reversible with respect to a distribution or measure. However

## Theorem

*If distribution  $\pi$  is such that  $P$  is reversible with respect to it, then  $\pi$  is invariant for  $P$ .*

*Proof:* Fix  $x$ .

$$\pi(x) = \sum_y \pi(x)p(x, y) = \sum_y \pi(y)p(y, x)$$

## Examples

We first consider irreducible birth and death chains. If the chain has an invariant law  $\pi$ , then this must be a reversible measure for the process: For  $(X_n)_{n \geq 0}$  starting in equilibrium  $\pi$ , we have that for every  $A \subset \mathbb{N}$ ,  $P(X_n \in A)$  does not change. Thus

$$P(X_0 \in A, X_1 \notin A) = P(X_0 \notin A, X_1 \in A)$$

Apply this to  $A = [0, n]$  to get

$$\pi(n)p_{nn+1} = \pi(n+1)p_{n+1n}$$

Given our process only makes nearest neighbour moves, this is enough.

## Examples

Thus to see if a B and D chain is positive recurrent we have the following recipe. We use  $p_i$  for  $p_{ii+1}$  and  $q_i$  for  $p_{ii-1}$ .

- Put  $\nu(0) = 1$
- Recursively choose  $\nu(n) = \nu(n-1)p_{n-1}/q_n = \prod_{k=1}^n p_{k-1}/q_k$
- If  $\sum_k \nu(n) = M < \infty$ , then the chain is positive recurrent with invariant  $\pi(n) = \nu(n)/M$  ; otherwise it is not positive recurrent.

## Examples contd

Consider a “random walk” on a finite graph  $G = (V, E)$ , :  $P_{x,y} = \frac{1}{d_x}$  (where  $d_x$  = the degree of  $x$ ) if  $x$  and  $y$  are neighbours; otherwise  $P_{x,y} = 0$ . Then the measure  $\nu(x) = d_x$  satisfies the detailed balance and so (for a connected graph or equivalently an irreducible chain), the invariant distribution is  $\nu/M$  where  $M = \sum_x d_x$

# Reversibility

If a Markov chain is reversible with respect to probability distribution  $\pi$ , then if  $(X_n)_{n \geq 0}$  starts in equilibrium (i.e.  $X_n$  has law  $\pi \forall n$ ), then if we fix  $N < \infty$ , then  $(Y_n)_{0 \leq n \leq N}$  given by  $Y_n = X_{N-n}$  satisfies for any  $i_0, i_1, \dots, i_N$

$$P(Y_0 = i_0, Y_1 = i_1 \dots Y_N = i_N) = P(X_0 = i_N, X_1 = i_{N-1} \dots X_N = i_0)$$

$$\pi(i_N) p_{i_N, i_{N-1}} \dots p_{i_1 i_0}$$

But using now reversibility we see

$$\begin{aligned} p_{i_{N-1}, i_N} \pi(i_{N-1}) p_{i_{N-1}, i_{N-2}} \dots p_{i_1 i_0} &= p_{i_{N-1}, i_N} p_{i_{N-2}, i_{N-1}} \pi(i_{N-2}) \dots p_{i_1 i_0} \dots \\ &= \pi(i_0) p_{i_0 i_1} p_{i_1 i_2} \dots p_{i_{N-1}, i_N}, \end{aligned}$$

# Reversibility

## Theorem

*If  $(X_n)_{n \geq 0}$  is a Markov chain in starting with an invariant probability with respect to which it is reversible, then for each  $N$*

$$(X_n)_{0 \leq n \leq N} \stackrel{D}{=} (X_{N-n})_{0 \leq n \leq N}$$

## An invariant measure.

Up until now we have really been interested in invariant distributions rather than invariant measures. Recall a measure on  $I$ ,  $\nu$  is simply a collection of positive values  $\nu(i) i \in I$ . If  $\sum_i \nu(i) < \infty$ , then we may divide  $\nu$  by its total mass to obtain a probability distribution. In later studies of continuous time process we will need some results on invariant measures, even if they are not finite.

# An invariant measure.

## Theorem

For an irreducible recurrent Markov chain fix  $k \in I$  and define

$$\gamma(j) = \mathbb{E}_k \left( \sum_{r < T_k} I_{X_r=j} \right)$$

Then

- $\gamma(k) = 1$
- $0 < \gamma(j) < \infty$
- $\gamma$  is invariant.

Proof (i) is immediate. Secondly, since the chain is irreducible for any  $j$ ,  $\mathbb{P}_k(T_j < T_k) > 0$  and  $\mathbb{P}_j(T_k < T_j) > 0$ . The first inequality implies that  $\gamma(j) > 0$  while the second one implies that under  $\mathbb{P}_k$ , the number of visits to  $j$  before  $T_k$  is stochastically dominated by a Geometric random variable of parameter  $\mathbb{P}_j(T_k < T_j) > 0$  and so  $\gamma(j) < \infty$

## An invariant measure.

For (iii), just as in the proof of the existence of the invariant probability, the key element is a temporal shift. Since the chain is recurrent we have  $T_k$  is a.s. finite. Thus

$$\gamma(j) = \mathbb{E}_k \left( \sum_{0 < r \leq T_k} I_{X_r=j} \right)$$

Now as before, we can write

$$\begin{aligned} \gamma(j) &= \mathbb{E}_k \left( \sum_{0 \leq r < T_k} I_{X_r=i} \sum_i p_{ij} \right) \\ &= \sum_i \nu(i) p_{ij} \end{aligned}$$

### Theorem

If  $\lambda$  is another invariant measure satisfying (i)-(iii) above, then  $\gamma = \lambda$ .

## Proof of Theorem.

We first prove  $\gamma \leq \lambda$ . Fix  $j \neq k$ , then

$$\lambda(j) = \sum_{i_1} \lambda(i_1) p_{i_1 j} = \lambda(k) p_{kj} + \sum_{i_1 \neq k} \lambda(i_1) p_{i_1 j}$$

$$= p_{kj} + \sum_{i_1 \neq k} \lambda(i_1) p_{i_1 j}$$

$$\sum_{i_2} \sum_{i_1 \neq k} \lambda(i_2) p_{i_2 i_1} p_{i_1 j} + p_{kj} = \sum_{i_2 \neq k} \sum_{i_1 \neq k} \lambda(i_2) p_{i_2 i_1} p_{i_1 j} + \sum_{i_1 \neq k} p_{ki_1} p_{i_1 j} + p_{kj}$$

Continuing

$$= \sum_{i_3} \sum_{i_2 \neq k} \sum_{i_1 \neq k} \lambda(i_3) p_{i_3 i_2} p_{i_2 i_1} p_{i_1 j} + \sum_{i_1 \neq k} p_{ki_1} p_{i_1 j} + p_{kj}$$

$$= \sum_{i_3 \neq k} \sum_{i_2 \neq k} \sum_{i_1 \neq k} \lambda(i_3) p_{i_3 i_2} p_{i_2 i_1} p_{i_1 j} + \sum_{i_2 \neq k} \sum_{i_1 \neq k} p_{k i_2} p_{i_2 i_1} p_{i_1 j} + \sum_{i_1 \neq k} p_{k i_1} p_{i_1 j} + p_{kj}$$

Continuing to arbitrary many iterations and noting that all terms are positive we obtain the inequality  $\lambda(j) \geq \gamma(j)$ .

To complete the argument we consider invariant measure  $\lambda - \gamma$ . This is positive everywhere and zero at  $k$ . But by iteration of the invariance, we must have the difference is zero for every  $j$  such that  $p_{jk} > 0$ , iterating again we see  $\lambda(j) - \gamma(j)$  must be zero for every  $j$  with  $p_{jk}^2 > 0$  and so on. By irreducibility we have every  $j$  must give value zero.

# Ergodic Theorem

## Theorem

For a positive recurrent irreducible Markov chain  $(X_n)_{n \geq 0}$  and a bounded  $f$  on  $I$

$$\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{k < n} f(X_k) \rightarrow \sum_{i \in I} \pi_i f(i)$$

a.s. over all initial distributions.

The proof is easy. we suppose that  $f$  is bounded by 1 in absolute value. We already know that for any  $k$ , the proportion of time spent in state  $k$  converges a.s. to  $\pi_k$ , so we get this law of large numbers for all  $k$  in any fixed finite subset of  $I$ . Let us fix  $A \subset I$  finite. Then  $\frac{1}{n} \sum_{k < n} f(X_k) =$

$$\sum_i N_n(i) f(i) = \sum_{i \in A} N_n(i) f(i) + \sum_{i \in A^c} N_n(i) f(i)$$

where  $N_n(i)$  is the proportion of time spent in state  $i$  by time  $n$

## proof contd

But  $\sum_{i \in A} N_n(i)f(i)$  converges to  $\sum_{i \in A} \pi_n f(i)$ . Recalling that  $f$  is bounded by one we get

$$\limsup_n \left| \frac{1}{n} \sum_{k < n} f(X_k) - \sum_{i \in I} \pi_i f(i) \right| \leq 2 \sum_{i \in A^c} \pi_i$$

## Generalization

We simply observe that for any integer  $r \geq 1$  (under the previous conditions)  $(Y_n)_{n \geq 0}$  is an irreducible Markov chain on the set of “possible”  $rt$ uples  $(i_1, i_2, \dots, i_r) \in I^r$  for  $Y_n = (X_n, X_{n+1}, \dots, X_{n+r-1})$ . It has unique invariant distribution

$$\pi^Y((i_1, i_2, \dots, i_r)) = \pi(i_1)p_{i_1, i_2} \cdots p_{i_{r-1}, i_r}.$$

### Theorem

For a positive recurrent irreducible Markov chain  $(X_n)_{n \geq 0}$ , positive integer  $r$  and a bounded  $f$  on  $I^r$

$$\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{k < n} f(X_k, X_{k+1}, \dots, X_{k+r-1}) \rightarrow \sum_{\bar{i} \in I^r} f(\bar{i}) \pi(i_1)p_{i_1, i_2} \cdots p_{i_{r-1}, i_r}$$

a.s. over all initial distributions.